Using the **Reconcile Algorithm**, we address discrepancies between **Conditional Average Treatment Effect (CATE) Estimators** to solve the **reference class problem** in causal inference for more consistent individual predictions.

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Background $|\equiv$

- leads to varied predictions, from limited data
- *●* aggregate performance but differ in individual predictions
- *●* Reconcile Algorithm: reconciles models to solve for

Research objectives

- \bullet Explain equivalence between predictive multiplicity and reference class problem
- Solve reference class problem in causal inference by applying the

predictive multiplicity, similar to multicalibration algorithm

Reconcile Algorithm to reconcile **CATE Estimators**

The Reference Class Problem across domains

Philosophy - Reference Class Problem

(Cosmos Club anecdote)

Definition. Reference Class Problem. In a universe X with a distribution over elements e and a probabilistic predicate $P : \mathcal{X} \to \{0, 1\}$, the individual reference class problem arises for e if there are distinct reference classes S_i and S_j such that (1) $e \in S_i$ and $e \in S_j$, (2) $S_i \neq S_j$, and (3) $S_i \nsubseteq S_j$, $S_j \nsubseteq S_i$. Furthermore, if $Pr[P(e)|e \in S_i] \neq Pr[P(e)|e \in S_i]$, this results in ambiguity in assigning a definitive probability to $P(e)$ based on a single reference class.

Statistics - Heterogeneous Effects

• Source of stereotyping and bias Under heterogeneous treatment effects,

 $Y(1)_i - Y(0)_i \neq Y(1)_j - Y(0)_j$

Reconciling Heterogeneous Effects in Causal Inference

predict individualized outcome:

 $CATE(X = x) = \mathbb{E}[Y(1) - Y(0)|X = x]$

On what basis to choose X?

Computer Science - Predictive Multiplicity

Reconcile Algorithm reconciles individual probability predictions

Definition. Predictive Multiplicity. Given a baseline classifier h_0 and an error tolerance ϵ , a prediction problem exhibits predictive multiplicity over the ϵ -level set $S_{\epsilon}(h_0)$ if there exists a classifier $h \in S_{\epsilon}(h_0)$ such that $h(x_i) \neq$ $h_0(x_i)$ for some instance x_i in the training dataset, indicating that competing models within the set achieve similar aggregate performance while differing in their individual predictions.

Dawid's Insight

Two approaches to individual prediction:

Individual to Group (i2G)

- Model individual-level predictions before aggregating to groups
- Calculate group probabilities by averaging within a \bullet reference class and compare to empirical averages to tune/falsify predictions
- Non-unique models \rightarrow predictive multiplicity problem!

Group to Individual (G2i)

• Start with aggregate data to derive individual predictions by selecting a sufficiently large reference class and using its proportion for estimates.

Reconcile Algorithm

Roth, Tolbert, and Weinstein propose the **Reconcile Algorithm** to solve predictive multiplicity / reference class problem.

- 1. Contest a model f_A with another model f_B if they disagree substantially on individual predictions. Extract a large reference class S from their disagreement region. At least one of the models has a lower mean squared error, and thus falsifies the other model.
- 2. **Update** the falsified model (WLOG) f_A to produce a new model f'_A that makes predictions that are correct on average over S. f'_{4} is now not falsified and more accurate.
- 3. Repeat steps 1-2 until f_A and f_B agree within some error bound.

The algorithm can be applied to any model f .

We apply Reconcile to address CATE estimator disagreement, adopting the i2G perspective: begin modeling individual treatment effects, rather than starting from group-level averages, and falsify to group-level statistics!

Expanded for clarity:

-
- How to select reference class? \rightarrow reference class problem
	- Inputs may belong to multiple reference classes; can't condition on all of them
	- Different choices of reference class lead to \circ different estimates

Equivalence!

- Both originate from data failing to encode unique ● estimates for individual probabilities
	- **G2i** data samples make it difficult after conditioning, curse of dimensionality
	- **i2G** challenging to generalize reference classes to make confident individual predictions
- Equivalent goal and limitation: data evidences multiple possible "true probabilities"
- No longer *choosing* ref class (choosing between CATE ●estimators) but reconciling CATE estimators via subgroup performance

A simple reduction

Construct a new variable, estimated treatment effect:

 $\hat{y} = \mathbb{E}[y \mid X = x, T = 1] - \mathbb{E}[y \mid X = x, T = 0].$

This makes our outcome continuous instead of discrete (necessary for Reconcile). Then, define the loss function over which Reconcile will minimize. We use Brier loss.

$$
B(\hat{\tau}, \mathcal{D}) = \mathbb{E}_{(x,y,t)\sim\mathcal{D}}\left[(\hat{\tau}(x) - (\mathbb{E}[y|X=x, T=1] - \mathbb{E}[y|X=x, T=0]))^2 \right]
$$

Hence we'll use the following group-level statistic to falsify our predictions:

$$
\frac{1}{n}\sum_{i=1}^n\left(\hat{\tau}(x_i)-\hat{y}_i\right)^2
$$

Note that this is equivalent to the Expected Mean Squared Error, a commonly-used metric for assessing CATE performance. So applying Reconcile can yield a unified model that optimally estimates the CATE.

```
Algorithm 1: ReconcileCATE
1 Let t = t_1 = t_2 = 0 and \hat{\tau}_1^{t_1} = \hat{\tau}_1, \hat{\tau}_2^{t_2} = \hat{\tau}_22 Let m = \left| \frac{2}{\sqrt{\alpha \epsilon}} \right|3 while \mu(U_{\epsilon}(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2})) \geq \alpha do
  4 | for each \bullet \in \{>, <\} and i \in \{1,2\} do
                        v_*^{\bullet} = \mathop{\mathbb{E}}_{(x,y,t \sim D)}[y|x \in U_{\epsilon}^{\bullet}(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2}), T = 1] - \mathop{\mathbb{E}}_{(x,y,t \sim D)}[y|x \in U_{\epsilon}^{\bullet}(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2}), T = 0]v_i^{\bullet} = \mathop{\mathbb{E}}_{(x,y,t\sim D)} [\hat{\tau}_i(x)^{t_i}|x\in U_{\epsilon}^{\bullet}(\hat{\tau}_1^{t_1},\hat{\tau}_2^{t_2})].\overline{7}Let(i_t, \bullet_t) = i \in \{1, 2\}, \bullet \in \{>, <\}\mu(U_{\epsilon}(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2})) \cdot (v_*^{\bullet} - v_i^{\bullet})^28<sup>1</sup>and
                                                                                         g_t(x) = \begin{cases} 1 & x \in U^{\bullet_t}(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2}) \ 0 & \text{otherwise} \end{cases}\boldsymbol{9}Let
                  \tilde{\Delta}_t = \left( \mathop{\mathbb{E}}_{(x,u,t)\sim \mathcal{D}}[y|g_t(x) = 1, T = 1] - \mathop{\mathbb{E}}_{(x,u,t)\sim \mathcal{D}}[y|g_t(x) = 1, T = 0] \right) - \mathop{\mathbb{E}}_{(x,u,t)\sim \mathcal{D}}[\hat{\tau}_{i_t}^{t_i}(x)|g_t(x) = 1]10
                                                                                                        \Delta_t = \text{Round}(\tilde{\Delta}_t, m)
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Let $\hat{\tau}_i^{t_i+1}(x) = h(x, \hat{\tau}_i^{t_i}, g_t, \Delta_t), t_i = t_i + 1, t = t + 1$ 11 Output $(\hat{\tau}_1^{t_1}, \hat{\tau}_2^{t_2})$. 12

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